

Missing data: best practice and beyond in flexible modelling and causal inference

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Overview

- ▶ Handling missing data: a perspective of the state of the art
- ▶ Using multiple imputation for missing data in non-linear & hierarchical models
- ▶ Missing data in marginal structural models
 - ▶ Illustrative example
 - ▶ Common methods used and their assumptions
 - ▶ Simulation study
- ▶ Discussion

A perspective on the state of the art

- ▶ Rubin published his classification of missing data mechanisms in 1976 [1], and his classic book on multiple imputation for surveys in 1987 [2].
- ▶ There are two algorithms for multiple imputation of missing data: *joint modelling (JM)* (c.f. [3]) and *full conditional specification (FCS)* (c.f. [4, 5]). Joint modelling is more natural for multilevel/hierarchical structures, and FCS for cross-sectional data, involving a mix of variable types (e.g. interval censored variables) and questionnaire features such as skips.
- ▶ Either FCS or JM (and often both) are now implemented in all standard software packages.

Challenges for practitioners

- ▶ The key challenge for practitioners is choosing an appropriate imputation model.
- ▶ This needs to be consistent with the scientific model. Analysts also need to choose which auxiliary variables, not in the scientific model, to additionally include in the imputation model.
- ▶ The STRATOS missing data topic group has developed guidance (<https://arxiv.org/abs/2004.14066>); see also [6], and STRATOS workshop at August 2020 ISCB.

Further challenge: handling non-linear relationships

A further challenge is how to handle non-linear relationships in the multiple imputation, particularly if combined with multilevel structure, e.g. for observations i on units j :

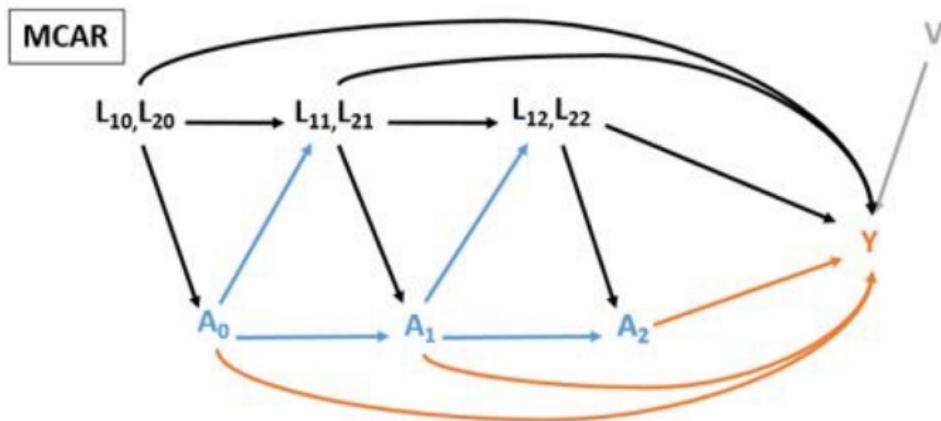
$$Y_{i,j} = (\beta_0 + u_{0,j}) + (\beta_1 u_{1,j})x_{1,i,j} + \beta_2 x_{2,i,j} + \beta_3 x_{2,i,j}^2 + \epsilon_{i,j}$$
$$\begin{pmatrix} u_{0,j} \\ u_{1,j} \end{pmatrix} \stackrel{iid}{\sim} N(\mathbf{0}, \Sigma)$$
$$\epsilon_{i,j} \stackrel{iid}{\sim} N(0, \sigma_e^2)$$

Solutions

- ▶ In single-level (cross sectional) models, one approach with FCS is to include all interactions in the imputation models [7]; however, this can get cumbersome, and is not always appropriate.
- ▶ A theoretically preferable approach is to construct imputations consistent with the non-linear structure [8]; this is available as `smcfc`s in R and Stata.
- ▶ Building on the approach proposed by [9], this has now been implemented in the R package for multilevel modelling, `jomo` [10], as `smcjomo`.

Missing data in Marginal Structural Models (MSMs)

MSMs were developed by Robins and co-workers, to estimate intervention effects from observational data affected by time varying confounding, for example:



where Y is the continuous outcome, A_0, A_1, A_2 time-varying binary treatments, L 's are time varying confounders (one binary, one continuous) and V is an additional variable predictive of outcome.

The challenge

- ▶ Estimation follows a two-stage process:
 1. weights — based on the inverse of the probability of a patient receiving the treatment they actually received — are estimated to create a pseudo-population in which treatment and confounders are independent.
 2. a weighted regression (using the weights derived in the first stage) including only the treatment history can be used to obtain estimate the causal effect of the treatment regimens of interest.
- ▶ In practice, the weights can be estimated using pooled logistic regression, in which each person-time interval is considered as an observation. This pooled logistic regression model must include the confounders and their relevant interactions to ensure the distributions of confounders are balanced between treatment groups
- ▶ However, there is no consensus on the appropriate method to use when the confounder data have a non-monotone missingness pattern.

Five commonly used missing data methods for partially observed confounders

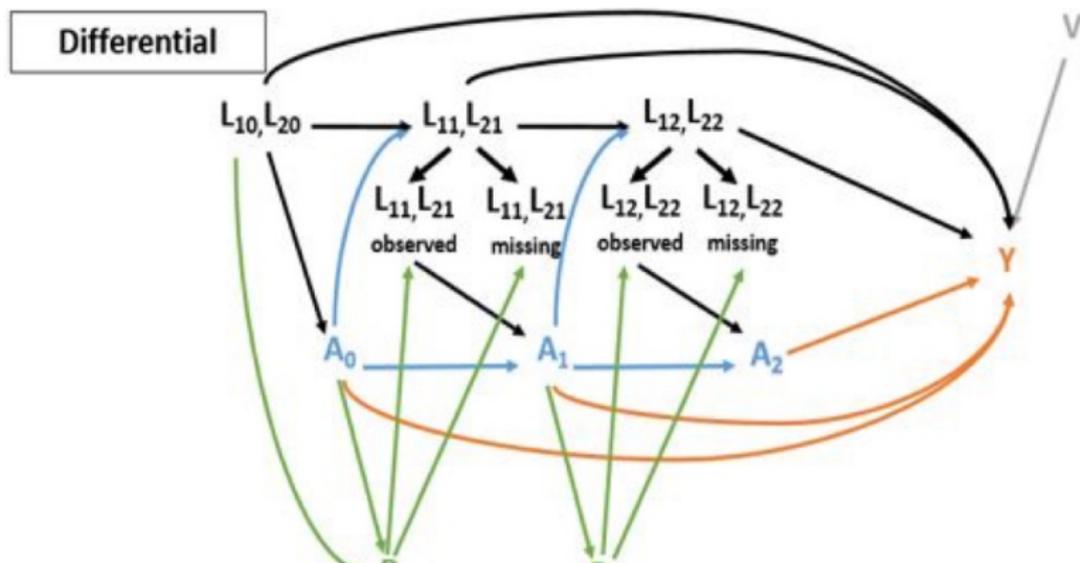
Method	Missing data on ...	Assumptions	Unbiased in MSMs when...	Advantages	Limitations
Complete case (CC)	Covariates Treatment Outcome	Missing data are MCAR	MCAR	Straightforward	May be inefficient because of the loss in sample size
Last Observation Carried Forward (LOCF)	Covariates Treatment Outcome except at baseline	The true, but missing, value is the same as the last available measurement OR The treatment decision depends on the previous available measurement rather than the true (unobserved) one	Constant	Straightforward Discards fewer patients from the analysis than CC	It can lead to too narrow confidence intervals Patients are discarded if baseline measurements are missing

...continued

Multiple imputation (MI)	Covariates Treatment Outcome	Missing data are <u>MAR^a</u> The imputation model is correctly specified	MCAR MAR <u>A,L</u> MAR A,L,Y MAR A,L,V	Maintains the original sample size	May be computationally intensive Challenging for a <u>large number of time points</u>
Inverse-probability-of-missingness weighting (IPMW)	Covariates Treatment Outcome	Missing data are MAR given the treatment and the covariates, but not the outcome The weight model is correctly specified	MCAR MAR <u>A,L</u> MAR A,L,V Constant	Faster than MI for large datasets Weights simultaneously address confounding and missing data	May be inefficient for small and moderate sample size
Missingness Pattern Approach (MPA)	Covariates	The partially observed covariate is no longer a confounder once missing e.g. the treatment decision depends on the confounder value only when a measurement is available	Differential	Relatively simple to implement Assumptions do not relate to Rubin's taxonomy so may work when standard methods do not	Does not handle missing data on the exposure or outcome Challenging when the number of missingness patterns is large

Details of simulation study

- ▶ We simulated data from $n = 10,000$ individuals with about 40% missing data in the confounders, and used 5000 replications.
- ▶ Values were informed by a motivating study of sleep apnoea. Full details in the supplementary materials of the forthcoming paper [11].



Results

True MSM is:

$$Y_i = \beta_{int} + 1.163a_{0,i} + 1.677a_{1,i} + 2a_{2,i} + \epsilon_i; \epsilon_i \stackrel{iid}{\sim} N(0, \sigma^2)$$

Absolute bias and coverage rate (%) for the 5 methods to handle missing data in each scenario considered at time 2:

Scenario	CC		LOCF		MPA		MI		IPMW	
	Bias	Coverage	Bias	Coverage	Bias	Coverage	Bias	Coverage	Bias	Coverage
MCAR	0.000	98.1	0.100	80.2	0.094	82.8	0.002	97.9	0.000	97.0
MAR AL	0.002	98.1	0.122	71.3	0.115	75.6	0.004	97.7	0.002	97.2
MAR ALY	-0.547	0.0	0.000	98.7	-0.437	0.0	0.002	98.3	-0.663	0.0
MAR ALV	-0.002	97.9	0.104	79.3	0.103	79.7	0.001	97.8	-0.003	96.8
Constant	0.001	98.1	0.001	97.8	0.165	49.9	0.095	83.8	0.001	97.6
Differential	-0.004	97.9	0.034	96.2	0.001	96.8	-0.048	94.9	-0.003	96.9

Full details in [11]

Summary & Discussion

- ▶ Multiple imputation provides a very general, applicable, method for handling missing data. It is particularly useful with missing covariates.
- ▶ The most common challenges are making imputation models consistent with the substantive model, and choosing appropriate auxiliary variables. The former can be addressed using `smcfc`s (Stata and R) and `jomo` and `smcjomo` in R.
- ▶ In MSMs, a variety of proposals for handling missing data have been made; we summarised them and reviewed their assumptions.
- ▶ Our results show that:
 - ▶ It is important to reflect carefully on the likely missing data mechanisms. If they assumptions of one of the similar methods really hold, this is preferable
 - ▶ MI had the best across-the-board performance, and is always worth doing, at least as a secondary analysis.
 - ▶ Better coverage could be obtained by accounting for weight estimation (e.g., with MI, [12]).

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